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To cite this version:
Jean-Bernard Lasserre. ON THE CHRISTOFFEL FUNCTION AND CLASSIFICATION IN DATA ANALYSIS. Comptes Rendus. Mathématique, 2022, 360, pp.919–928. 10.5802/crmath.358. hal-03620965v2

HAL Id: hal-03620965
https://laas.hal.science/hal-03620965v2
Submitted on 9 Dec 2022

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Volume 360 (2022), p. 919-928

https://doi.org/10.5802/crmath.358

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On the Christoffel function and classification in data analysis

Jean B. Lasserre

Abstract. We show that the empirical Christoffel function associated with a cloud of finitely many points sampled from a distribution, can provide a simple tool for supervised classification in data analysis, with good generalization properties.

Résumé. Nous montrons que la fonction de Christoffel empirique associée à un échantillon fini de points peut fournir un outil simple pour la classification supervisée en analyse de données, avec de bonnes propriétés de généralisation.

2020 Mathematics Subject Classification. 41A30, 42C05, 47B32, 68T09, 94A16.

Funding. J.B. Lasserre is supported by the AI Interdisciplinary Institute ANITI funding through the french program “Investing for the Future PIA3” under the grant agreement number ANR-19-PI3A-0004.

Manuscript received 16 March 2022, revised and accepted 21 March 2022.

1. Introduction

In this note we are mainly concerned with supervised classification with noiseless deterministic labels where the objects of interest \( x \in X \) belong to \( m \) classes with supports \( X_j \subseteq X \subseteq \mathbb{R}^n \), \( j \in [m] \) (with \( [m] = \{1, \ldots, m\} =: Y \)). The supports satisfy \( X_i \cap X_j = \emptyset \) for all \( i, j \) with \( i \neq j \). The data set consists of clouds of finitely many points \( \{x(i)\} \subseteq X_j \) sampled from an underlying distribution \( \phi_j \) on \( X_j \), \( j \in [m] \). In this situation, an exact classifier \( f : X \to Y \), selects \( j =: f(x) \) whenever \( x \in X_j \). When constructing a classifier from a sample of data points, as e.g. in machine learning, a sensitive issue is its generalization properties when applied on a test set different from the training set. For the reader interested in recent developments on various techniques and issues in supervised and unsupervised classification, we refer to e.g. the book [1] and the many references therein.
**Contribution.** We first introduce a simple and natural ideal classifier \( f_t : \mathbf{X} \to \mathbf{Y} \), with nice asymptotic properties as \( t \) increases. It is based on the Christoffel function \( \Lambda^\phi_t \) associated with the joint distribution \( d\mu(x,y) \) on \( \mathbf{X} \times \mathbf{Y} \). As \( \mu \) is supported on the graph \{ \( (x,f(x)) : x \in \mathbf{X} \) \} of the exact classifier \( f \), recent results of [5] transported in our context suggest that the classifier \( f_t(x) := \arg\max_{y \in \mathbf{Y}} \Lambda^\phi_t(x,y) \) should approximate \( f \) nicely. This is the case and indeed, by a slight modification of the definition of \( \Lambda^\phi_t \), we show that \( f_t \) is simply expressed in terms of the Christoffel functions \( \Lambda^\phi_t \) of the \( \phi_j \); namely, \( f_t(x) = \arg\max_k \Lambda^\phi_t(x) \). Notice that this simple form of \( f_t \) mathematically justifies for supervised classification, the intuitive argument that \( \Lambda^\phi_t(x) > \Lambda^\phi_k(x), \forall k \neq j \), whenever \( t \) is sufficiently large and \( x \in \mathbf{X}_j \). Indeed as \( x \in \mathbf{X}_j \) is outside the support \( \mathbf{X}_k \) of \( \phi_k \), for every \( k \neq j \), the “score” \( \Lambda^\phi_t(x) \) is close to zero for sufficiently large \( t \), as it decreases exponentially fast to zero (while the decrease of \( \Lambda^\phi_t(x) \) is at most polynomial in \( t \)).

We next consider the practical case where we only have access to a discrete sample of points in each class \( \mathbf{X}_j \) (e.g., the training set in Machine Learning) so that \( \Lambda^\phi_t \) is not available. We provide a data-driven analogue of the previous result which, as expected, is in terms of the Christoffel functions \( \Lambda^\phi_{j,N} \) associated with the discrete empirical measures \( \phi_j \). Namely the empirical discrete analogue \( f^N_t \) of the classifier \( f_t \) simply reads \( f^N_t(x) = \arg\max_k \Lambda^\phi_{j,N}(x) \), and has same properties as \( f_t \), but of course in an almost-sure sense with respect to random samples. In particular it shows good generalization properties. Indeed with \( \varepsilon > 0 \) fixed and \( t \) sufficiently large, with probability 1 (with respect to random samples), \( f^N_t(x) = j \) for every \( j \in [m] \) and all \( x \in \mathbf{X}_j \) at distance at least \( \varepsilon \) from the boundary \( \partial \mathbf{X}_j \), for sufficiently large \( N \).

Finally, we also briefly discuss more general joint distributions of pairs \((x,y)\) (where possibly \( \mathbf{X}_i \cap \mathbf{X}_j \neq \emptyset \) for some \((i,j)\)) which covers practical cases where some misclassification may occur and/or some ambiguity is allowed.

1.1. Notation, definitions and preliminary results

Let \( \mathbb{R}[x] \) denote the ring of real polynomials in the variables \( x = (x_1, \ldots, x_n) \) and \( \mathbb{R}[x]_t \subset \mathbb{R}[x] \) be its subset of polynomials of total degree at most \( t \). Let \( \mathbb{N}^t_\mathbb{N} := \{\alpha \in \mathbb{N}^n : |\alpha| \leq t \} \) (where \( |\alpha| = \sum_i \alpha_i \) with cardinal \( s(t) = \binom{n+t}{n} \). Let \( \Phi_t(x) = (x^\alpha)_{\alpha \in \mathbb{N}^t_\mathbb{N}} \) be the vector of monomials up to degree \( t \).

The support of a Borel measure \( \mu \) on \( \mathbb{R}^n \) is the smallest closed set \( A \) such that \( \mu(\mathbb{R}^n \setminus A) = 0 \), and such a set \( A \) is unique.

**Moment matrix.** Let \( \phi \) be a Borel measure whose support \( \Omega \subset \mathbb{R}^n \) is compact with nonempty interior. Its moment matrix of order (or degree) \( t \) is the real symmetric matrix \( \mathbf{M}_t(\phi) \) with rows and columns indexed by \( \mathbb{N}^t_\mathbb{N} \), and with entries

\[
\mathbf{M}_t(\phi)(\alpha,\beta) := \int_\Omega x^{\alpha+\beta} \, d\phi = \phi_{\alpha+\beta}, \quad \alpha,\beta \in \mathbb{N}^t_\mathbb{N}.
\]

Then necessarily \( \mathbf{M}_t \) is positive semidefinite for all \( t \), denoted \( \mathbf{M}_t(\phi) \succeq 0 \).

**Christoffel function.** If \( \Omega \) has nonempty interior then \( \mathbf{M}_t \) is positive definite for all \( t \), denoted \( \mathbf{M}_t(\phi) > 0 \). Let \( (P_\alpha)_{\alpha \in \mathbb{N}^n} \subset \mathbb{R}[x] \) be a family of polynomials, orthonormal with respect to \( \phi \), i.e.,

\[
\int_\Omega P_\alpha P_\beta \, d\phi = \delta_{\alpha=\beta}, \quad \forall \alpha,\beta \in \mathbb{N}^n.
\]

Then the Christoffel function (CF) \( \Lambda^\phi_t : \mathbb{R}^n \to \mathbb{R}_+ \) associated with \( \phi \), is defined by

\[
x \mapsto \Lambda^\phi_t(x) := \left[ \sum_{\alpha \in \mathbb{N}^n} P_\alpha(x)^2 \right]^{-1}, \quad \forall x \in \mathbb{R}^n,
\]

(1)
and recalling that $M_t(\mu)$ is nonsingular, it turns out that
\[ \Lambda_t^\phi(x) = \left[ v_t(x)^T M_t(\phi)^{-1} v_t(x) \right]^{-1}, \quad \forall x \in \mathbb{R}^n. \] (2)
An equivalent and variational definition is also
\[ \Lambda_t^\phi(x) = \inf_{p \in \mathbb{R}[x]} \left\{ \int_\Omega p^2 \, d\phi : p(x) = 1 \right\}, \quad \forall x \in \mathbb{R}^n. \] (3)
One interesting and distinguishing feature of the CF is that as $t$ increases, $\Lambda_t^\phi(x)$ ↓ 0 exponentially fast for every $x \notin \text{support}(\phi)$. In other words $\Lambda_t^\phi$ identifies the support of $\phi$ when $t$ is sufficiently large. In addition, at least in dimension $n = 2$ or $n = 3$, one may visualize this property even for small $t$, as the resulting superlevel sets $\Omega_\gamma := \{ x : \Lambda_t^\phi(x) \geq \gamma \}, \gamma \in \mathbb{R}$, capture the shape of $\Omega$ quite well; see e.g. [2].

1.2. Setting

Let $Y := \{1, 2, \ldots, m\}$ be the set of $m$ classes, and for each (class) $j \in Y$, let $X_j \subset \mathbb{R}^n$ be the set of points in the class $j$, assumed to be open with compact closure $\overline{X}_j$. Let $X := \bigcup_{j=1}^m X_j \subset \mathbb{R}^n$ be the open set (with compact closure $\overline{X} = \bigcup_{j=1}^m \overline{X}_j$) of all points to be classified and $\mu$ be the joint probability distribution of $(x, y)$ on $\overline{X} \times Y$. Write
\[ d\mu(x, y) = \varphi(dy|x) \phi(dx), \] (4)
where $\mu$ has been disintegrated into its marginal $\phi$ on $\overline{X}$ and its conditional probability distribution $\varphi(dy|x)$ on $Y$ given $x \in \overline{X}$. Next, each point $x \in X$ belongs to only one class and therefore $X_i \cap X_j = \emptyset$ for all pairs $(i, j)$ with $i \neq j$, and so we may and will assume that $\varphi(\overline{X}_i \cap \overline{X}_j) = 0$ for all pairs $(i, j)$ with $i \neq j$. Therefore one may write $\mu = \sum_{j=1}^m \mu_j$, with
\[ d\mu_j(x, y) = \delta_{ij}(dy) \varphi_j(dx), \quad j \in Y, \]
for some marginals $\varphi_j$ on $\overline{X}_j, j \in Y$. In particular:
\[ \phi(A) = \mu(A \times Y) = \sum_{j=1}^m \mu_j(A \times Y) = \sum_{j=1}^m \varphi_j(A), \quad \forall A \in \mathcal{B}(X), \]
and therefore $\phi = \sum_{j=1}^m \varphi_j$. Next, let $f : \overline{X} \to Y$ be the exact classifier
\[ x \mapsto f(x) = \begin{cases} \sum_{j=1}^m j \cdot 1_{X_j}(x) & \text{if } x \in X, \\ 0 & \text{otherwise}. \end{cases} \] (5)
So $f(x)$ identifies the class of $x \in X$ and returns $f(x) = 0$ if $x$ belong to some intersection $\overline{X}_i \cap \overline{X}_j$. Notice that one may write
\[ d\mu(x, y) = \delta_{\infty}(dy) \phi(dx), \] (6)
and so the joint distribution $\mu$ is supported on the graph $G := \{(x, f(x)) : x \in \overline{X}\}$ of the function $f$.

The Christoffel function is a powerful tool from the theory of approximation and orthogonal polynomials and one of its distinguishing features is its ability to identify the support of the underlying measure. So the CF $\Lambda_t^\mu$ associated with $\mu$ is an appropriate tool to approximate $f$ since the graph of $f$ is precisely the support of the measure $\mu$ in (6). In Marx et al. [5] the authors propose to approximate $f$ (when $\|f\|_{\infty} < M$) by:
\[ \hat{f}_t(x) := \arg\min_y \Lambda_t^{\mu+\epsilon\mu_0}(x, y)^{-1}, \quad \forall x \in X, \] (7)
with a small $\epsilon > 0$ and where $\mu_0$ is a measure with a density w.r.t. Lebesgue measure, positive on $X \times [-M, M]$. They prove nice theoretical convergence guarantees as $t$ increases; see [5] for more details. Notice that in the present supervised classification framework, the function to
approximate is a step function so that the support of its graph is contained in a real algebraic variety, an even more specific case.

2. Main result

We first consider the ideal case of approximating the classifier \( f \) in (5) via the CF of the joint distribution \( \mu \) in (6). Then we next consider the more practical setting (as in machine learning) where we only have access to a finite sample (the training set). In this case we use the empirical measures \( \phi_{j,N} \) associated with the points of the sample in class “\( j \)”. In Section 2.2 we invoke results from [4] that relate the degree \( t \) of the CF \( \Lambda_{t}^{\phi_{j,N}} \) with the size \( N \) of the sample to ensure that important asymptotic properties of \( \Lambda_{t}^{\phi_{j,N}} \) and \( \Lambda_{t}^{\phi_{j}} \) as \( t \) and \( N \) increase, coincide.

2.1. The CF on a real variety

Let \( q \in \mathbb{R}[x,y] \) be the polynomial \( (x,y) \mapsto v(x,y) = \prod_{i=1}^{n}(y-i) \) and let \( \mu \) be the probability measure on \( \Omega = \mathbb{X} \times Y \) defined in (6). Its support \( \Omega \) is contained in the real algebraic variety \( V := \{ (x,y) : v(x,y) = 0 \} \subset \mathbb{R}^{n} \times Y \) and the ideal \( \mathcal{I} = \langle v \rangle \subset \mathbb{R}[x,y] \) generated by the polynomial \( v \) is the ideal of polynomials that vanish on \( V \).

Observe that the moment matrix \( M_{t}(\mu) \) is singular since the vector \( v \) of coefficients of the polynomial \( v \) is in the kernel of \( M_{t}(\mu) \) as soon as \( t \geq m \). Indeed \( v^{T}M_{t}(\mu)v = \int \nu^{2}d\mu = 0 \) because the support of \( \mu \) is contained in \( V \). So the definition (2) of the CF is not valid anymore. Denote by \( L_{t}^{2}(\mu) \subset \mathbb{R}[x,y] \) the space of polynomials on \( V \) of total degree at most \( t \) (and degree at most \( m-1 \) in the variable \( y \)), equipped with the inner product and norm inherited from \( L^{2}(\Omega, \mu) \). It turns out that \( L_{t}^{2}(\mu) \) is a RKHS (Reproducing Kernel Hilbert Space). Then in this context, the variational definition (3) of the CF associated with \( \mu \) reads

\[
(x, y) \mapsto \Lambda_{t}^{\mu}(x, y) := \inf_{p \in L_{t}^{2}(\mu)} \left\{ \int_{\Omega} p^{2}d\mu : p(x,y) = 1 \right\}, \quad \forall (x,y) \in V.
\]

The set \( \Gamma_{t} := \{ x^{\alpha}y^{k} : k \leq m-1; |\alpha| + k \leq t \} \subset \mathbb{R}[x,y] \) is a monomial basis of \( L_{t}^{2}(\mu) \). Let \( M'_{t}(\mu) \) be the moment matrix associated with \( \mu \) in (6) with rows and columns indexed by all monomials \( (x^{\alpha}y^{k}) \) of \( \Gamma_{t} \) (and not all monomials \( x^{\alpha}y^{k} \) of total degree at most \( t \)), e.g., listed according to the lexicographic ordering. Then \( M'_{t}(\mu) \) is non singular and

\[
\Lambda_{t}^{\mu}(x, y)^{-1} = v'_{t}(x, y)^{T}M'_{t}(\mu)^{-1}v'_{t}(x, y), \quad \forall (x,y) \in V,
\]

where \( v'_{t}(x, y) \) is the vector of all monomials of \( \Gamma_{t} \). Alternatively

\[
\Lambda_{t}^{\mu}(x, y)^{-1} = \sum_{(\alpha,k) \in \Gamma_{t}} \frac{1}{\lambda_{\alpha,k}} Q_{\alpha,k}(x,y)^{2},
\]

where the \( Q_{\alpha,k} \)'s and the \( \lambda_{\alpha,k} \)'s are the eigenvectors and their respective eigenvalues associated with \( M'_{t}(\mu) \). Following [5], one may consider the perturbed measure \( \mu + \epsilon \mu_{0} \) where \( \mu_{0} \) is a probability uniformly distributed on \( \mathbb{X} \times [0,m] \) and \( \epsilon > 0 \) is a small parameter. Then

\[
\Lambda_{t}^{\mu+\epsilon \mu_{0}}(x, y)^{-1} = v_{t}(x, y)^{T}M_{t}(\mu+\epsilon \mu_{0})^{-1}v_{t}(x, y), \quad \forall (x,y) \in \mathbb{R}^{n} \times \mathbb{R}.
\]

Recall that \( v_{t}(x, y) \) is the vector of all monomials \( x^{\alpha}y^{k} \) of degree at most \( t \), and \( M_{t}(\mu+\epsilon \mu_{0}) \) is the moment matrix of \( \mu + \epsilon \mu_{0} \) of order \( t \), which is non singular for all \( \epsilon > 0 \). In addition \( \Lambda_{t}^{\mu+\epsilon \mu_{0}} \) is defined for \( \forall (x,y) \in \mathbb{R}^{n+1} \) whereas \( \Lambda_{t}^{\mu} \) in (8) is defined for \( \forall (x,y) \in V \). Then as proved in [5] the classifier \( \hat{f}_{t} \) in (7) approximates \( f \) as \( t \) increases. Next we show that by a slightly change of the vector space \( L_{t}^{2}(\mu) \) in (8), the resulting CF has nice additional properties that can be exploited to provide the resulting classifier (7) with a clear and more intuitive interpretation.
A slight variant of the CF. We now introduce a slight variant $\hat{\Lambda}_\mu^t$ of the CF $\Lambda_\mu^t$, defined by:

\[
(x, y) \mapsto \hat{\Lambda}_\mu^t(x, y) := \inf_{p \in \mathcal{L}_t^2(\mu)} \left\{ \int_{\mathcal{O}} p^2 \, d\mu : p(x, y) = 1 \right\}, \quad \forall (x, y) \in V, \tag{9}
\]

where $\mathcal{L}_t^2(\mu) = \mathbb{R}[x, y]_{t, m-1}$ is the vector space of polynomials of degree at most $t$ with respect to the variable $x$ and at most $m-1$ with respect to the variable $y$, so that $L_t^2(\mu) \subset \mathcal{L}_t^2(\mu) \subset L_t^2(m-1, \mu)$.

**Proposition 1.** Let $\Lambda_\mu^t$ and $\hat{\Lambda}_\mu^t$ be as in (8) and (9) respectively. Then:

\[
\Lambda_{i+m-1}^t(x, y) \leq \hat{\Lambda}_{i+m-1}^t(x, y) \leq \Lambda_{i}^t(x, y), \quad \forall (x, y) \in V. \tag{10}
\]

**Proof.** Follows from $L_t^2(\mu) \subset \mathcal{L}_t^2(\mu) \subset L_t^2(m-1, \mu)$ and the definitions (8) and (9). 

Proposition 1 states that $\hat{\Lambda}_\mu^t$ and $\Lambda_\mu^t$ are close but as we next see, $\hat{\Lambda}_\mu^t$ has an interesting additional feature. Namely, it has a nice characterization in closed form which when exploited for classification leads to a classifier with a clear interpretation. Let $(\theta_j)_{j \in \mathbb{N}} \subset \mathbb{R}[y]_{m-1}$ be the interpolation polynomials at the points $\{1, 2, \ldots, m\}$ of $Y$, i.e.,

\[
y \mapsto \theta_j(y) := \prod_{i \neq j} (y - i) / \prod_{i \neq j} (j - i), \quad i = 1, \ldots, m,
\]

which form an orthonormal family with respect to the uniform probability measure on $Y$.

**Theorem 2.** For each $j \in Y$, let $(P_\alpha^j)_{\alpha \in \mathbb{N}_0^n} \subset \mathbb{R}[x]$ be a family of polynomials that are orthonormal with respect to the marginal probability measure $\phi_j$ of $\mu_j$, and let $\Lambda_\mu^j$ be the standard Christoffel function associated with $\phi_j$ on $\mathbb{R}_+$. Then:

(i) The family $(\theta_j(y) P_\alpha^j(x))_{\alpha \in \mathbb{N}_0^n} \subset \mathbb{R}[x, y]$ is an orthonormal basis of $\mathcal{L}_t^2(\mu)$.

(ii) The Christoffel function $\hat{\Lambda}_\mu^t$ defined in (9) satisfies

\[
\hat{\Lambda}_\mu^t(x, y)^{-1} = \sum_{j \in Y} \theta_j(y)^2 \sum_{\alpha \in \mathbb{N}_0^n} P_\alpha^j(x)^2, \quad \forall (x, y) \in \mathbb{R}^n \times Y
\]

\[
= \sum_{j \in Y} \theta_j(y)^2 \Lambda_\mu^j(x)^{-1}, \quad \forall (x, y) \in \mathbb{R}^n \times Y \tag{11}
\]

\[
= \sum_{j \in Y} \delta_{y=j} \Lambda_\mu^j(x)^{-1}, \quad \forall (x, y) \in \mathbb{R}^n \times Y. \tag{12}
\]

**Proof.** (i). Every element of $\mathcal{L}_t^2(\mu) = \mathbb{R}[x, y]_{t, m-1}$ is of the form $\sum_{j=0}^{m-1} y^j q_j(x)$ with $q_j \in \mathbb{R}[x]_t$. Hence consider a polynomial $u \in \mathcal{L}_t^2(\mu)$ in the form $u(x, y) := p(y) q(x)$ for some $q \in \mathbb{R}[x]_t$ and some $p \in \mathbb{R}[y]_{m-1}$, arbitrary. Then as $(P_\alpha^j)_{\alpha \in \mathbb{N}_0^n}$ generates $\mathbb{R}[x]_t$, observe that for every $j \in Y:

\[
q(x) = \sum_{\alpha \in \mathbb{N}_0^n} q_\alpha^j P_\alpha^j(x), \quad \forall x \in \mathbb{R}^n,
\]

for some coefficients $(q_\alpha^j)_{\alpha \in \mathbb{N}_0^n}$. Next, as the polynomials $(\theta_j)_{j \in Y}$ generate $\mathbb{R}[y]_{m-1}$, write $p(y) = \sum_{j \in Y} p_j \theta_j(y)$ for some coefficients $(p_j)_{j \in Y}$, and therefore

\[
u(x, y) = \sum_{j \in Y} p_j (\theta_j(y) q(x)) = \sum_{\alpha \in \mathbb{N}_0^n, j \in Y} p_j q_\alpha^j \left[ \theta_j(y) P_\alpha^j(x) \right].
\]

**Orthogonality.** If $i \neq j$ then $\theta_i(y) \theta_j(y) = 0$ everywhere on the support of $\mu$ and therefore

\[
\int_{\mathcal{O}} \theta_i(y) P_\alpha^i(x) \theta_j(y) P_\beta^j(x) \, d\mu(x, y) = 0,
\]
whereas if \( i = j \) then
\[
\int_{\Omega} \theta_i(y)^2 P_i^j(\mathbf{x}) P_j^j(\mathbf{x}) \, d\mu(\mathbf{x}, y) = \sum_{j=1}^{m} \int_{\Omega} \theta_i(y)^2 P_i^j(\mathbf{x}) P_j^j(\mathbf{x}) \, d\mu_j(\mathbf{x}, y)
\]
\[
= \int_{\Omega} \theta_i(y)^2 P_i^j(\mathbf{x}) P_j^j(\mathbf{x}) \, d\mu_j(\mathbf{x}, y) = \int_{\Omega} P_i^j(\mathbf{x}) P_j^j(\mathbf{x}) \, d\phi_i(\mathbf{x}) = \delta_{\alpha=\beta}.
\]

Next, as \( \mathcal{L}_t^2(\mu) \subset \mathbb{R}[\mathbf{x}, y]_{t,m-1} \), its cardinality is \( r(t) := m \cdot \binom{n+t}{t} \) which is also the number of terms in the family \( \{ \theta_j(y) P_a^j(\mathbf{x}) \}_{a \in \Omega_r, j \in \mathcal{Y}} \) which also generates \( \mathcal{L}_t^2(\mu) \). Hence \( \{(\theta_j(y) P_a^j(\mathbf{x}))_{a \in \Omega_r, j \in \mathcal{Y}} \) is an orthonormal basis of \( \mathcal{L}_t^2(\mu) \).

(ii). Let \( \overline{M}_t(\mu) \) be the moment matrix of degree \( t \) with rows and columns indexed by monomials \( (x^a y^b)_{a \in \Omega_r, 0 \leq j \leq m-1} \) (e.g. with lexicographic ordering), and let \( \tilde{v}_t(\mathbf{x}, y) \) be the vector of monomials \( x^a y^b \) listed with the same ordering. Observe that (9) reads
\[
\hat{\Lambda}^\mu_t(\mathbf{x}, y) = \min \left\{ \mathbf{p}^T \hat{M}_t(\mu) \mathbf{p} : \langle \mathbf{p}, \tilde{v}_t(\mathbf{x}, y) \rangle = 1 \right\},
\]
where \( \mathbf{p} \in \mathbb{R}^{r(t)} \) is the vector of coefficients of \( p \in \mathcal{L}_t^2(\mu) \) in that basis. Then (14) is a convex optimization problem whose optimal solution \( \mathbf{p}^* \) satisfies \( 2 \hat{M}_t(\mu) \mathbf{p}^* = \lambda^* \tilde{v}_t(\mathbf{x}, y) \) for some scalar \( \lambda^* \). Hence \( \lambda^* = 2 \hat{\Lambda}^\mu_t(\mathbf{x}, y) \) and
\[
\mathbf{p}^* = \hat{\Lambda}^\mu_t(\mathbf{x}, y) \hat{M}_t(\mu)^{-1} \tilde{v}_t(\mathbf{x}, y)
\]
so that the corresponding polynomial \( p^* \in \mathcal{L}_t^2(\mu) \) reads
\[
p^*(\mathbf{u}, z) = \hat{\Lambda}^\mu_t(\mathbf{x}, y) \tilde{v}_t(\mathbf{u}, z)^T \hat{M}_t(\mu)^{-1} \tilde{v}_t(\mathbf{x}, y)
\]
\[
= \hat{\Lambda}^\mu_t(\mathbf{x}, y) \sum_{a \in \Omega_r, j \in \mathcal{Y}} \theta_j(y) \theta_j(z) P_a^j(\mathbf{u}) P_a^j(\mathbf{x}), \quad (\mathbf{u}, z) \in \mathbb{R}^n \times \mathcal{Y},
\]
and therefore
\[
1 = p^*(\mathbf{x}, y) = \hat{\Lambda}^\mu_t(\mathbf{x}, y) \sum_{a \in \Omega_r, j \in \mathcal{Y}} \theta_j(y)^2 P_a^j(\mathbf{x})^2, \quad \forall (\mathbf{x}, y) \in \mathbb{R}^n \times \mathcal{Y},
\]
which is (11). In particular we also retrieve that
\[
\hat{\Lambda}^\mu_t(\mathbf{x}, y)^{-1} = \tilde{v}_t(\mathbf{x}, y)^T \hat{M}_t(\mu)^{-1} \tilde{v}_t(\mathbf{x}, y), \quad \forall (\mathbf{x}, y) \in \mathcal{V}.
\]
Next (12) follows from the definition of the Christoffel function associated with \( \phi_j \) for each \( j \in \mathcal{Y} \), and (13) follows from the properties of interpolation polynomials \( (\theta_j)_{j \in \mathcal{Y}} \).

So whenever \( y \in \mathcal{Y} \), the Christoffel function \( \hat{\Lambda}^\mu_t(\mathbf{x}, y) \) has a very simple expression (12), stated directly in terms of the Christoffel functions \( (\Lambda^\mu_j(\mathbf{x}))_{j=1, \ldots, m} \) associated with the classes \( j = 1, \ldots, m \). This is quite natural but is proper to the CF \( \hat{\Lambda}^\mu_t \) and not to the standard CF \( \Lambda^\mu_t \).

An ideal classifier. Given the Christoffel function \( \hat{\Lambda}^\mu_t \) defined in (12) and inspired by (7), a natural candidate classifier is the function
\[
\mathbf{x} \mapsto \hat{f}_t(\mathbf{x}) := \arg \min_{y \in \mathcal{Y}} \hat{\Lambda}^\mu_t(\mathbf{x}, y)^{-1} \arg \max_{j \in \mathcal{Y}} \hat{\Lambda}^\mu_t(\mathbf{x}, y), \quad \forall \mathbf{x} \in \mathcal{X},
\]
which in view of (13) reads:
\[
\mathbf{x} \mapsto \hat{f}_t(\mathbf{x}) := \max_{k \in [m]} \Lambda^\phi_k(\mathbf{x}), \quad \forall \mathbf{x} \in \mathcal{X}.
\]
Observe that the “max” in (17) is over \( y \in [m] \) and not over the interval \([0, m]\). This is because in supervised classification, we know that \( f(\mathbf{x}) \in [m] \) for all \( \mathbf{x} \in \mathcal{X} \). The rationale is the following: Let \( \mathbf{x} \in \mathcal{X}_j \) be fixed arbitrary, so that \( \mathbf{x} \notin \mathcal{X}_k \) for every \( k \neq j \). As \( t \) increases, \( \Lambda^\phi_k(\mathbf{x}) \) decreases to
zero exponentially while \( \Lambda_t^\phi_j(x) \) decreases not faster than \( t^n \). Therefore for \( t \) sufficiently large, necessarily \( \Lambda_t^\phi_k(x) < \Lambda_t^\phi_j(x) \) for all \( k \neq j \), and so by (13),

\[
x \in X_j \Rightarrow \exists \ t_0 \ \text{s.t.} \ \hat{f}_t(x) = \arg\min_{y \in Y} \Lambda_t^\mu(x, y)^{-1} = \arg\max_k \Lambda_t^\phi_k(x) = j, \quad \forall \ t \geq t_0.
\]

An even stronger almost-uniform result holds. Let \( \partial X_j \) denote the boundary of the set \( X_j \subset \mathbb{R}^n \).

**Theorem 3.** Let \( \hat{f} \) be as in (17) and let \( X_j^\varepsilon := \{ x \in X_j : d(x, \partial X_j) > \varepsilon \} \) where \( \varepsilon > 0 \) is fixed. Assume that for every \( j \in [m] \), \( \phi_j \) has a density w.r.t. Lebesgue measure \( \lambda \) restricted to \( X_j \), bounded from below by \( c > 0 \). Then there exists \( t_\varepsilon \) such that \( \hat{f}_t(x) = j \) for all \( x \in X_j^\varepsilon \) and all \( t \geq t_\varepsilon \).

**Proof.** Let \( s(t) := \binom{n+t}{t} \) and denote by \( \text{diam}(S) \) the diameter of a bounded set \( S \subset \mathbb{R}^n \). Let \( \gamma_k := \phi_k(X_k), \ k \in [m] \). If \( x \in X_j \) then \( d(x, \overline{X}_k) > \varepsilon \) for all \( k \neq j \). Hence by [3, Lemma 6.6] (and using that \( \Lambda_t^{\phi_k/\gamma_k} = \Lambda_t^{\phi_k} / \gamma_k \)),

\[
\frac{\gamma_k}{s(t)} \Lambda_t^{\phi_k}(x)^{-1} \geq 2^{\text{diam}(X_k)-3} t^{-n} (\frac{n}{e})^n \exp(-n^2/t), \quad \forall \ x \in X_j^\varepsilon.
\]

On the other hand, as \( d(x, \partial X_j) > \varepsilon \), we can invoke [3, Lemma 6.2] extended to measures with density \( \rho \) Lebesgue bounded from below by \( e > 0 \) (see [3, Assumption 3.11]) and use \( (t+1)(t+2)(t+3)/(n+1)(n+2)(n+2t+6) \geq 1/2(n+1)^3 \), to obtain

\[
\frac{\gamma_j}{s(t)} \Lambda_t^{\phi_j}(x)^{-1} \leq \frac{2 \lambda(X_j)}{e \varepsilon^n \omega_n} (1 + n)^3, \quad \forall \ x \in X_j^\varepsilon,
\]

where \( \omega_n \) is the \( n \)-dimensional area of \( S^{n+1} \). Hence clearly there exists \( t_\varepsilon \) such that \( \Lambda_t^{\phi_k}(x)^{-1} \geq \Lambda_t^{\phi_j}(x)^{-1} \) for all \( k \neq j \) and all \( x \in X_j^\varepsilon \), whenever \( t \geq t_\varepsilon \). In particular, as the functions \( \Lambda_t^{\phi_k} \) are strictly positive, continuous and \( \overline{X}_j \) is compact, there exists \( a_j > 0 \) such that

\[
\forall \ x \in X_j^\varepsilon : \ \Lambda_t^{\phi_k}(x) < \Lambda_t^{\phi_j}(x) - a_j, \quad \forall \ k \neq j.
\]

Therefore the result follows from (13) and (16). \( \square \)

### 2.2. Application to supervised classification with noiseless deterministic labels

In supervised classification we do not have access to the CF \( \Lambda_t^\mu \) or \( \hat{\Lambda}_t^\mu \). We only have access to a sample of \( N \) points \( \text{Tr}_N = \{(x(i), y(i)) : i = 1, \ldots, N \} \subset X \) (the training data set) and a sample of test points (the test data set). For instance, in a typical Machine Learning (ML) approach one tries to learn a classifier function \( f \) in (5) from the supervised data \( \text{Tr}_N \) by computing parameters of a deep neural network that minimize some loss function. Usually, the number of parameters is very large (compared to \( N \)) making the resulting solution sensitive to a classical overfitting phenomenon. One way to attenuate this overfitting phenomenon is to add an appropriate regularization term to the loss function in the criterion to minimize.

One reason behind this overfitting phenomenon is that in minimizing the loss function, each data point \((x(i), y(i))\) is treated separately. Ideally one should somehow consider the entire training set \( \text{Tr}_N \) itself and not its members separately. This is precisely what the CF function approach does. Indeed the training set \( \text{Tr}_N \) is used to construct the empirical (discrete) analogues \( \phi_{j,N} \) of the measures \( \phi_j \), and their associated empirical Christoffel function \( \Lambda_t^{\phi_{j,N}} \), now obtained from empirical moments. Remarkably, even though the geometry of the support of \( \phi_{j,N} \) is quite trivial, the CF \( \Lambda_t^{\phi_{j,N}} \) is still close to \( \Lambda_t^{\phi_j} \) in a certain sense and the training set \( \text{Tr}_N \) can still be used to infer properties of the underlying measures \( \phi_j \). Hence, and importantly, even though the mathematical object \( \Lambda_t^{\phi_{j,N}} \) is built from individual items, it is in fact concerned with the cloud of points of \( \text{Tr}_N \) in class \( \{j\} \), rather than the points \( x(i) \) in that class taken separately. However
of course, for $\Lambda_{t}^{N}$ to recover asymptotic properties of $\Lambda^{N}$, the sample size $N$ and the degree $t$ cannot be chosen independently; see [3, 4] for more details.

**Setting.** For every $j \in [m]$, let $\mathbf{Tr}_{N}^{j} = (\mathbf{x}(i))_{i \leq N} \subset \mathbf{X}_{j}$ be a training set for class $\{j\}$, where $\mathbf{x}(i)$ are i.i.d. random vectors with common distribution $\phi_{j}$ whose support is $\mathbf{X}_{j}$. So the whole training set $\mathbf{Tr}_{N}$ has a total of $mN$ points where the $N$ points in each class $j$ are sampled from $\phi_{j}$. For every fixed $t$, a natural approach suggested by (17), consists of:

- Computing the Christoffel function $\Lambda_{t}^{N}$ associated with the empirical prob. measure

$$\phi_{j,N} := \frac{1}{N} \sum_{x(i) \in \mathbf{X}_{j}} \delta_{\mathbf{x}(i)}, \quad \forall \; j \in [m]. \tag{19}$$

Following (2), $\Lambda_{t}^{N}^{-1}(\mathbf{x}) = v_{t}(\mathbf{x})^{T} \mathbf{M}_{t}(\phi_{j,N})^{-1} v_{t}(\mathbf{x})$, for all $\mathbf{x} \in \mathbb{R}^{n}$ and all $j \in [m]$. The moments of $\phi_{j,N}$ are easily obtained by

$$\phi_{j,N}(\alpha) := \left\{ \frac{1}{N} \sum_{i} \mathbf{x}(i)^{A} : \mathbf{x}(i) \in \mathbf{Tr}_{N}^{j} \right\}, \quad \forall \; \alpha \in \mathbb{N}^{n},$$

and the moment matrix $\mathbf{M}_{t}(\phi_{j,N})$ is nonsingular for sufficiently large $t$.

- Following (17), introduce the empirical classifier

$$\mathbf{x} \rightarrow \tilde{f}_{N}^{j}(\mathbf{x}) := \arg \max_{k \; \in \; \mathbb{X}_{j}} \Lambda_{t}^{N}(\mathbf{x}), \quad \forall \; \mathbf{x} \in \mathbb{X}_{j}. \tag{20}$$

Then the empirical version of Theorem 3 reads as follows:

**Theorem 4.** For every $j \in [m]$, let $(\mathbf{x}(i))_{i \leq N} \subset \mathbf{X}_{j}$ be i.i.d. random vectors according to a distribution $\phi_{j}$ whose support is $\mathbf{X}_{j}$ and which satisfies the assumption in Theorem 3. Let $\phi_{j,N}$ be as in (19), and $\tilde{f}_{N}^{j}$ be as in (20). Given $\varepsilon > 0$ fixed, let $\mathbf{X}_{j}^{\varepsilon} := \{ \mathbf{x} \in \mathbf{X}_{j} : d(\mathbf{x}, \partial \mathbf{X}_{j}) > \varepsilon \}$, $j \in [m]$. Then there exists $t_{\varepsilon}$ such that for all $t > t_{\varepsilon}$ fixed, with probability 1 (with respect to the random samples $\mathbf{Tr}_{N}^{k} \subset \mathbf{X}_{k}$, $k \in [m]$), $\tilde{f}_{N}^{N}(\mathbf{x}) = j$ for all $\mathbf{x} \in \mathbf{X}_{j}^{\varepsilon}$, for sufficiently large $N$.

**Proof.** By Theorem 3 there exist $t_{\varepsilon}$ and $\delta_{\varepsilon} > 0$, such that $\Lambda_{t}^{N}(\mathbf{x}) - \Lambda_{t}^{N}(\mathbf{x}) > \delta_{\varepsilon}$ for all $k \neq j$, $\forall \; t \geq t_{\varepsilon}$ and all $\mathbf{x} \in \mathbf{X}_{j}^{\varepsilon}$, see (18). On the other hand, with $t$ fixed, by [3, Theorem 3.13],

$$\sup_{\mathbf{x} \in \mathbb{R}^{n}} \left| \Lambda_{t}^{N}(\mathbf{x}) - \Lambda_{t}^{N}(\mathbf{x}) \right|_{a.s.} \rightarrow 0, \quad \text{as } N \rightarrow \infty,$$

where the “a.s.” is with respect to the random sample $\mathbf{Tr}_{N}^{k} \subset \mathbf{X}_{j}$. Hence for every $t > t_{\varepsilon}$ fixed, with probability 1 (with respect to the $m$ random samples $\mathbf{Tr}_{N}^{k} \subset \mathbf{X}_{k}$, $k \in [m]$)

$$\forall \; \mathbf{x} \in \mathbf{X}_{j}^{\varepsilon}, \forall \; k \neq j : \; \Lambda_{t}^{N}(\mathbf{x}) < \Lambda_{t}^{N}(\mathbf{x}) - \delta_{\varepsilon},$$

for sufficiently large $N$, and therefore $\tilde{f}_{N}^{N}(\mathbf{x}) = j$ for all $\mathbf{x} \in \mathbf{X}_{j}^{\varepsilon}$, $j \in [m]$. \hfill $\Box$

2.3. **The general case where the supports $\mathbf{X}_{j}$ are not disjoint**

We next briefly consider the case where the assumption $\mathbf{X}_{i} \cap \mathbf{X}_{j} = \emptyset$ for all $i \neq j$, is *not* satisfied. That is, misclassifications occur or some points $\mathbf{x} \in \mathbf{X}$ may indeed belong to several classes, as can be the case in some practical situations.

So let $\mu$ on $\bar{\mathbf{X}} \times \mathbf{Y}$, $\phi(\mathrm{d}y | \cdot)$ on $\mathbf{Y}$ and $\phi$ on $\bar{\mathbf{X}}$ be as in (4), but now with possibly $\mathbf{X}_{i} \cap \mathbf{X}_{j} = \emptyset$ for some $i \neq j$. For each $j \in [m]$, introduce the measures $\phi_{j}$ on $\mathbf{X}_{j}$, $j \in [m]$, defined by:

$$\phi_{j}(\mathrm{d}x) := \phi(j | x) \phi(\mathrm{d}x), \quad j \in \mathbf{Y}. \tag{21}$$

So in contrast to Section 2.1, $\phi_{j}(\mathbf{X}_{i}) \neq 0$ is allowed for $i \neq j$. Then an analogue of Theorem 2 reads:
Theorem 5. For each \( j \in Y \), let \((p_i^j)_{a \in \mathbb{N}} \subset \mathbb{R}[x] \) be a family of polynomials that are orthonormal with respect to the measure \( \phi_j \) on \( X \) defined in (21), and let \( \Lambda_i^{\phi_j} \) be the standard Christoffel function associated with \( \phi_j \). Then:

\( \Lambda_i^{\phi_j}(x) = \) where \( i \)

(i) The family \((\theta_i(y) P_i^{j}(x))_{a \in \mathbb{N}} \subset \mathbb{R}[x, y] \) is an orthonormal basis of \( L^2(\mu) \).

(ii) The Christoffel function \( \Lambda_i^{\phi_j} \) defined in (9) satisfies

\[
\Lambda_i^{\phi_j}(x) = \sum_{j \in \mathbb{N}} \theta_i(y) \sum_{a \in \mathbb{N}} P_i^{j}(x)^2, \quad \forall (x, y) \in \mathbb{R}^n \times Y
\]

\[
= \sum_{j \in \mathbb{N}} \theta_i(y)^2 \Lambda_i^{\phi_j}(x)^{-1} = \sum_{j \in \mathbb{N}} \delta_{y=j} \Lambda_i^{\phi_j}(x)^{-1}, \quad \forall (x, y) \in \mathbb{R}^n \times Y.
\]

Proof. (i). If \( i \neq j \) then \( \theta_i(y) \theta_j(y) = 0 \) everywhere on the support of \( \mu \) and therefore

\[
\int_{\Omega} \theta_i(y) P_i^{j}(x) \theta_j(y) P_i^{j}(x) \, d\mu(x, y) = 0,
\]

whereas if \( i = j \) then

\[
\int_{\Omega} \theta_i(y)^2 P_i^{j}(x) \, d\mu(x, y) = \int_{\Omega} P_i^{j}(x) \, d\mu(x, y) = \int_{\Omega} \theta_i(y)^2 \phi(dy|x) \, d\mu(x, y)
\]

\[
= \int_{\Omega} P_i^{j}(x) \, d\mu(x) = \delta_{\alpha=\beta}.
\]

(ii). The rest of the proof is similar to that of Theorem 2.

By (23), \( \Lambda_i^{\phi_j} \) is easily obtained as \( \Lambda_i^{\phi_j}(\cdot, j) \), \( j \in \mathbb{N} \), and \( \Lambda_i^{\phi_j}(x, y) \) is in turn obtained for instance via (15) from the moment matrix \( \mathbf{M}_t(\mu) \) of the joint distribution \( \mu \). As the CF is an appropriate tool for support inference, notice that intersections of super level sets \( G_{i,y} \cap G_{j,y} \), with \( G_{i,y} := \{ x : \Lambda_i^{\phi_j}(x) \geq \gamma \} \), should provide indications on whether \( X_i \cap X_j = \emptyset \) if \( i \neq j \).

Once again the CF \( \Lambda_i^{\phi_j} \) has the simple and nice expression (23) only in terms of the CF’s \( \Lambda_i^{\phi_j} \), which suggests to define a classifier \( \hat{f}_i(x) \) exactly as in (16). The only difference is the meaning of \( \phi_j \) and its implications. For instance if \( X_i \cap X_j \neq \emptyset \) then a point \( x \in X \) can belong to \( \text{supp}(\phi_i) \cap \text{supp}(\phi_j) \) with \( i \neq j \), and therefore the two scores \( \Lambda_i^{\phi_j}(x) \) and \( \Lambda_j^{\phi_j}(x) \) can be comparable even for large \( t \), whereas before for sufficiently large \( t \), the score \( \Lambda_i^{\phi_j}(x) \) (with \( j = \text{class}(x) \)) clearly dominates all other scores. In particular there is no analogue of Theorem 3. Evaluating how efficient the resulting classifier \( \hat{f}_i \) can be in a practical empirical context of sampled data as in Section 2.2, is beyond the scope of the present note.

3. Conclusion

The Christoffel function can provide a simple tool in supervised classification, with some theoretical guarantees. However to obtain \( \Lambda_i^{\phi_j,N} \) explicitly one must handle matrices of size \( (n \times t)^2 \) (inversion or eigenvectors) with a computational cost \( O(t^n) \) that grows rapidly with \( t \). Therefore so far, in this form this tool is limited to small dimension problems. On the other hand, evaluation of \( \Lambda_i^{\phi_j,N}(\xi) \) at a point \( \xi \in \mathbb{R}^n \) via (8) only requires to solve a simple convex quadratic optimization problem, which can be done efficiently even for large \( n \). Finally a detailed analysis of possible learning rates that can be obtained with this method remains to be done.
References


